

Poisson Process

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Outline

- Poisson Random Variable
- Non-Uniform Case
- Weiner Process
- White Noise Process
- Normal Process

Poisson Random Variable $X(\xi)$

$$P\{X(\xi) = k\} = e^{-a} \frac{a^k}{k!} \quad f_X(x) = \sum_{k=0}^{\infty} e^{-a} \frac{a^k}{k!} \delta(x - k)$$



$$E\{X\} = a$$

$$E\{X^2\} = a^2 + a$$

$$\sigma_x^2 = a$$

Consider probability experiment of placing points at random on a line.
Let $n(t_1, t_2)$ – number of points $\in (t_1, t_2)$;
 $X(t) = n(0, t)$ is a Poisson random variable if

- i) $P\{n(t_1, t_2) = k\} = e^{-\lambda t} \frac{(\lambda t)^k}{k!} \quad t = t_2 - t_1$
- ii) If intervals (t_1, t_2) & (t_3, t_4) are non-overlapping, random variables $n(t_1, t_2)$ & $n(t_3, t_4)$ are independent.

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X(t) is a Poisson process with parameter λt

$$P\{X(t) = k\} = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$E\{X(t)\} = \lambda t$$

$$E\{X^2(t)\} = \lambda^2 t^2 + \lambda t$$

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To obtain autocorrelation of X(t), assume $t_2 > t_1$ and consider $E\{X(t_1)[X(t_2) - X(t_1)]\} = E\{X(t_1)X(t_2)\} - E\{X^2(t_1)\}$

Noting that (0,t₁) and (t₁,t₂) do not overlap,

$$E\{X(t_1)\}E\{X(t_2) - X(t_1)\} = R(t_1, t_2) - (\lambda^2 t_1^2 + \lambda t_1)$$

or $\lambda t_1[\lambda(t_2 - t_1)] = R(t_1, t_2) - \lambda^2 t_1^2 - \lambda t_1$

$\Rightarrow R(t_1, t_2) = \lambda^2 t_1 t_2 + \lambda t_1 \quad t_2 \geq t_1$

Poisson Process- Autocorrelation Clarkson University

Similarly:

$$R(t_1, t_2) = \lambda^2 t_1 t_2 + \lambda t_2 \quad t_2 \geq t_1$$

$\Rightarrow R(t_1, t_2) = \lambda^2 t_1 t_2 + \lambda \min(t_1, t_2)$

Poisson Process- Non-Uniform Clarkson University

If the points on the line have non-uniform density $\lambda(t)$, λt must be replaced by $\int_0^t \lambda(\tau) d\tau$:

$$P\{X(t) = k\} = e^{-\int_0^t \lambda(\tau) d\tau} \frac{\left(\int_0^t \lambda(\tau) d\tau\right)^k}{k!} \quad E\{X(t)\} = \int_0^t \lambda(\tau) d\tau$$

$$R(t_1, t_2) = \int_0^{t_1} \lambda(\tau_1) d\tau_1 \int_0^{t_2} \lambda(\tau_2) d\tau_2 + \int_0^{\min(t_1, t_2)} \lambda(\tau) d\tau$$

Weiner Process Brownian Motion

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$W(t)$ is a Wiener Process (Brownian Motion) when:

i) $W(t)$ is a normal process with

$$\text{ii) } f(w;t) = \frac{1}{\sqrt{2\pi\alpha t}} e^{-\frac{w^2}{2\alpha t}} \quad F(w;t) = \frac{1}{2} + \text{erf} \frac{w}{\sqrt{\alpha t}}$$

ii) Independent increment process, i.e. $W(t_2) - W(t_1)$ is independent of $W(t_4) - W(t_3)$ if (t_1, t_2) and (t_3, t_4) are non-overlapping

iii) $W(0) = 0$

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Statistics of Wiener Processes

Mean $\Rightarrow E\{W(t)\} = 0$

Variance $\Rightarrow E\{W^2(t)\} = \alpha t$

Autocorrelation

$$R(t_1, t_2) = E\{W(t_1)W(t_2)\} = \begin{cases} \alpha t_2 & t_1 \geq t_2 \\ \alpha t_1 & t_2 \geq t_1 \end{cases} = \alpha \min(t_1, t_2)$$

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White Noise Process

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A White Noise Process is the derivative of a Wiener Process. That is,

$$n(t) = \frac{dW(t)}{dt}$$

Mean $\Rightarrow E\{n(t)\} = 0$

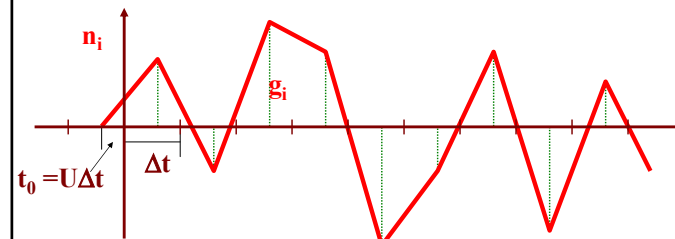
Autocorrelation $R(t_1, t_2) = \alpha \delta(t_1 - t_2)$

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Numerical Simulation of White Noise Process

- i) For a duration T (~ 20 s) select a small time step Δt (~ 0.01 to 0.05 s) and divide this into $m = T/\Delta t$ (~ 400 to 2000) subintervals
- ii) Generate $m+1$ zero-mean unit-variance normally distributed random numbers G_1, \dots, G_{m+1} . Multiply these by $\left(\frac{2\pi S_0}{\Delta t}\right)^{\frac{1}{2}}$ where S_0 is the constant power spectrum of the white noise. Evaluate $g_i = \left(\frac{2\pi S_0}{\Delta t}\right)^{\frac{1}{2}} G_i$

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iii) The white noise process is then given by

$n(t_0 + i\Delta t) = g_i, i = 1, 2, \dots, m+1, n(t_0) = 0$; n varies linearly over each subinterval. Here, t_0 is a random variable with uniform density over the subinterval $(-\Delta t, 0)$.

Transformation Form Pair of Uniform Random Variable to Gaussian

$$G_2 = \sqrt{-2 \ln U_1} \sin 2\pi U_2$$

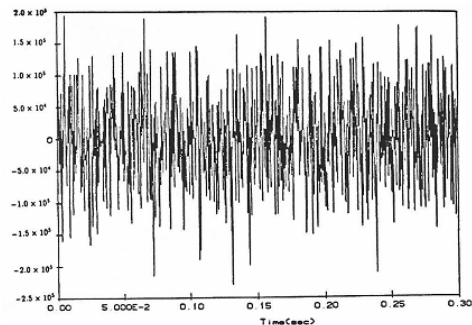
$$G_1 = \sqrt{-2 \ln U_1} \cos 2\pi U_2$$

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Sample White Noise Process, $d_p = 0.05 \mu\text{m}$

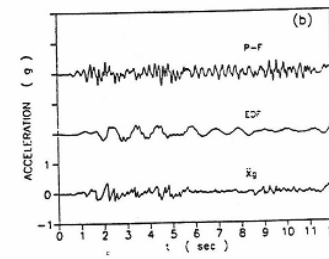
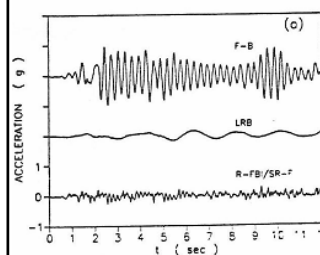


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Sample Responses Clarkson University

Sample Absolute Acceleration Responses at Top of Structure for El Centro 1940 Quake



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Normal Process $X(t)$ Clarkson University

If $X(t_1), X(t_2), \dots, X(t_n)$ jointly normal

$$f(x; t) = \frac{1}{\sqrt{2\pi C(t, t)}} e^{-\frac{(x-\eta(t))^2}{2C(t, t)}} \quad \text{(1st Order Density)}$$

(nth Order Density)

$$f(x_1, \dots, x_n; t_1, \dots, t_n) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Lambda|^{\frac{1}{2}}} \exp\left\{-\frac{1}{2} \sum_i \sum_j \Lambda_{ij}^{-1} (x_i - \eta(t_i))(x_j - \eta(t_j))\right\}$$

$$\Lambda = [C(t_i, t_j)] \quad |\Lambda| = \det|\Lambda| \quad \text{(Matrix of Covariance)}$$

(Note: Linear Combinations of normal processes are also normal)

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Concluding Remarks Clarkson University

- Poisson Random Variables
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Thank you!

Questions?

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